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# LASSO is not Fully Bayesian

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# LASSO is not fully Bayesian.

- Data Generation Model:  $Y = X\beta_0 + \epsilon$  where  $\epsilon \sim \mathcal{N}(0, I)$ .
- The LASSO optimization problem:

$$\hat{\beta}_\lambda^{\text{LASSO}} = \arg \min_{\beta \in \mathbb{R}^p} [\|Y - X\beta\|_2^2 + 2\lambda\|\beta\|_1] \quad (1)$$

- The posterior mode for the prior  $\beta_i \sim \text{Laplace}(\lambda)$ . Hence it has a Bayesian flavor.
- **Frequentist Optimality:** Can attain the (near) minimax rate  $O(s \log n)$  for the square Euclidean loss over  $s$ -sparse signals, if  $\lambda \approx \sqrt{2 \log(n)}$ .



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# Not really Bayesian!

- We will show that the full posterior does not contract at the same speed as its mode!
- Useless for uncertainty quantification, the central idea of Bayesian inference.
- The good behavior of the LASSO estimators must be due to the sparsity-inducing form of the posterior mode, **not** the Bayesian connection.



## Theorem (7) from [CSV15]

Assume that  $X = I$ . For any  $\lambda = \lambda_n$  such that  $\sqrt{n}/\lambda_n \rightarrow \infty$ , there exists  $\delta > 0$  such that, as  $n \rightarrow \infty$ ,

$$\mathbb{E}_{\beta^0=0} \Pi_{\lambda_n}^{\text{LASSO}} \left( \beta : \|\beta\|_2 \leq \delta \sqrt{n} \left( \frac{1}{\lambda_n} \wedge 1 \right) \mid Y \right) \rightarrow 0.$$

- Let  $\lambda_n = \sqrt{2 \log(n)}$ . The posterior places **no** weight on the ball  $\|\beta\|_2 = O\left(\sqrt{\frac{n}{\log n}}\right)$  which is substantially larger than the minimax rate  $\sqrt{s \log(n)}$  (unless the signal is extremely dense).



## Lemma (5.2) form [CV12]

We will first state two lemmas:

### Lemma (5.2) form [CV12]

For any prior probability distribution  $\Pi_n$  on  $\mathbb{R}^n$ , any positive measure  $\tilde{\Pi}_n$  with  $\tilde{\Pi}_n \leq \Pi_n$ , and any  $\beta_0$ ,

$$\int \frac{p_\beta}{p_{\beta_0}}(Y) d\Pi_n(\beta) \geq \|\tilde{\Pi}_n\| e^{-\tilde{\sigma}^2/2 + \tilde{\mu}^T(Y - \beta_0)},$$

where  $\tilde{\mu} = \int (\beta - \beta_0) d\tilde{\Pi}_n(\beta) / \|\tilde{\Pi}_n\|$  and  $\tilde{\sigma}^2 = \int \|\beta - \beta_0\|_2^2 d\tilde{\Pi}_n(\beta) / \|\tilde{\Pi}_n\|$ . Also, for any  $r > 0$ ,

$$P_{\beta_0} \left( \int \frac{p_\beta}{p_{\beta_0}}(Y) d\Pi_n(\beta) \geq e^{-r^2} \Pi_n(\beta : \|\beta - \beta_0\|_2 < r) \right) \geq 1 - e^{-r^2/8}.$$



## Proof of Lemma (5.2): First Assertion

We can write:

$$\begin{aligned}\log \left( \int \frac{p_\beta}{p_{\beta_0}}(Y) \frac{d\Pi_n(\beta)}{\|\tilde{\Pi}_n\|} \right) &\geq \int \log \left( \frac{p_\beta}{p_{\beta_0}}(Y) \right) \frac{d\Pi_n(\beta)}{\|\tilde{\Pi}_n\|} && \text{(Jensen's)} \\ &\geq \int \log \left( \frac{p_\beta}{p_{\beta_0}}(Y) \right) \frac{d\tilde{\Pi}_n(\beta)}{\|\tilde{\Pi}_n\|} && (\tilde{\Pi}_n \leq \Pi_n) \\ &= -\frac{\tilde{\sigma}^2}{2} + \tilde{\mu}^\top (Y - \beta_0) && \text{(Gaussianity)}\end{aligned}$$

Where  $\tilde{\mu} = \int (\beta - \beta_0) \frac{d\tilde{\Pi}_n}{\|\tilde{\Pi}_n\|}$  and  $\tilde{\sigma}^2 = \int \|\beta - \beta_0\|_2^2 \frac{d\tilde{\Pi}_n}{\|\tilde{\Pi}_n\|}$ .

Hence:

$$\int \frac{p_\beta}{p_{\beta_0}}(Y) d\Pi_n(\beta) \geq \|\tilde{\Pi}_n\| e^{-\tilde{\sigma}^2/2 + \tilde{\mu}^\top (Y - \beta_0)}.$$



## Proof of Lemma (5.2): Second Assertion

- Let  $\tilde{\Pi}_n$  be  $\Pi_n$  restricted to the set  $\{\beta : \|\beta - \beta_0\|_2 \leq r\}$ . Trivially  $\tilde{\Pi}_n \leq \Pi_n$ ,  $\|\tilde{\mu}\| \leq r$ ,  $\tilde{\sigma} \leq r$ , and  $\|\tilde{\Pi}_n\| = \Pi_n(\beta : \|\beta - \beta_0\|_2 < r)$ .
- Under  $P_{\beta_0}$ , we have that  $\tilde{\mu}^\top(Y - \beta_0)$  has the same distribution as  $\|\tilde{\mu}\|Z$  where  $Z \sim \mathcal{N}(0, I_n)$ .

- Note that

$$P\left[Zr \leq -r^2 + r^2/2\right] \leq \exp(-r^2/8).$$

- Hence:

$$P_{\beta_0}\left(\int \frac{p_\beta}{p_{\beta_0}}(Y) d\Pi_n(\beta) \geq e^{-r^2} \Pi_n(\beta : \|\beta - \beta_0\| < r)\right) \geq 1 - e^{-r^2/8}.$$

Which concludes the proof.  $\square$





## Lemma (7.1) from [CV12]

We have  $E_{\beta_0} \Pi_n(\beta : \|\beta - \beta_0\| < s_n | Y) \rightarrow 0$ , for any  $s_n$  for which there exist  $r_n$  such that

$$\frac{\Pi_n(\beta : \|\beta - \beta_0\| < s_n)}{\Pi_n(\beta : \|\beta - \beta_0\| < r_n)} = o(e^{-r_n^2}).$$



## Proof of Lemma (7.1)

Based on the previous lemma, there exists an event  $\mathcal{A}_n$  with  $\Pr[\mathcal{A}_n] \geq 1 - \exp(-r_n^2/8)$  such that

$$\int \frac{p_\beta}{p_{\beta_0}}(Y) d\Pi_n(\beta) \geq e^{-r^2} \Pi_n(\beta : \|\beta - \beta_0\| < r).$$

$$\begin{aligned} \mathbf{E}_{\beta_0} \left[ \Pi_n(\beta : \|\beta - \beta_0\| < s_n | Y) \mathbf{1}_{\mathcal{A}_n} \right] &= \mathbf{E}_{\beta_0} \left[ \frac{\int_{\beta: \|\beta - \beta_0\| < s_n} p_\beta(Y) d\Pi_n(\beta)}{\int p_\beta(Y) d\Pi_n(\beta)} \mathbf{1}_{\mathcal{A}_n} \right] \\ &\leq \mathbf{E}_{\beta_0} \left[ \frac{\int_{\beta: \|\beta - \beta_0\| < s_n} \frac{p_\beta}{p_{\beta_0}}(Y) d\Pi_n(\beta)}{e^{-r_n^2} \Pi_n(\beta : \|\beta - \beta_0\| \leq r_n)} \right] \\ &\leq \frac{\Pi_n(\beta : \|\beta - \beta_0\| < s_n)}{e^{-r_n^2} \Pi_n(\beta : \|\beta - \beta_0\| \leq r_n)} \rightarrow 0. \end{aligned}$$

Which concludes the proof. □



# Proof of the Main Theorem

Assume that  $\beta_0 = 0$ . We will use lemma (7.1) to prove this theorem. We will need to show that for proper sequences  $r_n$  and  $s_n$ ,

$$\frac{\Pi_n(\beta : \|\beta\|_2 < s_n)}{\Pi_n(\beta : \|\beta\|_2 < r_n)} = o(e^{-r_n^2}).$$

- Under the prior,  $|\beta_i| \sim \text{Exp}(\lambda)$  and  $\|\beta\|_1 \sim \text{Gamma}(\lambda, n)$ . Thus,

$$\begin{aligned}\Pi_n(\beta : \|\beta\|_2 < s_n) &\leq \Pi_n(\beta : \|\beta\|_1 < s_n \sqrt{n}) \\ &= \int_0^{\sqrt{ns_n}} \lambda^n u^{n-1} \frac{e^{-\lambda u}}{\Gamma(n)} du \leq \frac{(\lambda \sqrt{ns_n})^n}{\Gamma(n+1)}.\end{aligned}$$

- We can write,

$$\Pi_n(\beta : \|\beta\|_2 < r_n) = \left(\frac{\lambda}{2}\right)^n \int_{\|\beta\|_2 \leq r_n} e^{-\lambda \|\beta\|_1} d\beta \geq \left(\frac{\lambda}{2}\right)^n e^{-\lambda \sqrt{nr_n}} v_n r_n^n,$$

where  $v_n$  is the area of unit  $l_2$  ball in  $\mathbb{R}^n$  and the last inequality follows from  $\|\beta\|_1 \leq \sqrt{n} \|\beta\|_2$ .

By substituting  $v_n = \pi^{\frac{n}{2}}/\Gamma(1 + \frac{n}{2})$ , and last two inequalities,

$$\begin{aligned}\frac{\Pi(\|\beta\|_2 < s_n)}{\Pi(\|\beta\|_2 < r_n)} &\leq \left(\frac{2}{\sqrt{\pi}}\right)^n \frac{\Gamma(\frac{n}{2} + 1)}{\Gamma(n + 1)} n^{n/2} \exp\left(\lambda\sqrt{nr_n} - n \log\left(\frac{r_n}{s_n}\right)\right) \\ &\leq \exp\left(\lambda\sqrt{nr_n} - n \log\left(c\frac{r_n}{s_n}\right)\right).\end{aligned}$$

Set  $r_n = \sqrt{n}(\max(\lambda^{-1}, 1))$  and  $s_n = \delta r_n$ . Under these assumptions, the conditions of lemma (7.1) hold and we have:

$$\mathbb{E}_{\beta_0} \Pi_n(\beta : \|\beta - \beta_0\| < s_n | Y) \rightarrow 0,$$

equivalently,

$$\mathbb{E}_{\beta^0=0} \Pi_{\lambda_n}\left(\beta : \|\beta\|_2 \leq \delta\sqrt{n}\left(\frac{1}{\lambda_n} \wedge 1\right) \middle| Y\right) \rightarrow 0.$$

Which concludes the proof of the main theorem. □



- CSV15 Ismaël Castillo, Johannes Schmidt-Hieber, and Aad van der Vaart, Bayesian Linear Regression with Sparse Priors, *Annals of Statistics*, 2015.
- CV12 Ismaël Castillo, and Aad van der Vaart, Needles and Straw in a Haystack: Posterior Concentration for Possibly Sparse Sequences, *Annals of Statistics*, 2012.